## 1 Probability : Worked Examples

(1) The information entropy of a distribution $\left\{p_{n}\right\}$ is defined as $S=-\sum_{n} p_{n} \log _{2} p_{n}$, where $n$ ranges over all possible configurations of a given physical system and $p_{n}$ is the probability of the state $|n\rangle$. If there are $\Omega$ possible states and each state is equally likely, then $S=\log _{2} \Omega$, which is the usual dimensionless entropy in units of $\ln 2$.

Consider a normal deck of 52 distinct playing cards. A new deck always is prepared in the same order (A $2 \boldsymbol{\infty} \cdots \mathrm{~K}$ ).
(a) What is the information entropy of the distribution of new decks?
(b) What is the information entropy of a distribution of completely randomized decks?

Now consider what it means to shuffle the cards. In an ideal riffle shuffle, the deck is split and divided into two equal halves of 26 cards each. One then chooses at random whether to take a card from either half, until one runs through all the cards and a new order is established (see figure).
(c) What is the increase in information entropy for a distribution of new decks that each have been shuffled once?
(d) Assuming each subsequent shuffle results in the same entropy increase (i.e. neglecting redundancies), how many shuffles are


Figure 1: The riffle shuffle. necessary in order to completely randomize a deck?

Solution :
(a) Since each new deck arrives in the same order, we have $p_{1}=1$ while $p_{2, \ldots, 52!}=0$, so $S=0$.
(b) For completely randomized decks, $p_{n}=1 / \Omega$ with $n \in\{1, \ldots, \Omega\}$ and $\Omega=52$ !, the total number of possible configurations. Thus, $S_{\text {random }}=\log _{2} 52!=225.581$.
(c) After one riffle shuffle, there are $\Omega=\binom{52}{26}$ possible configurations. If all such configurations were equally likely, we would have $(\Delta S)_{\text {riffle }}=\log _{2}\binom{52}{26}=48.817$. However, they are not all equally likely. For example, the probability that we drop the entire left-half deck and then the entire right half-deck is $2^{-26}$. After the last card from the left half-deck is dropped, we have no more choices to make. On the other hand, the probability for the sequence LRLR $\cdots$ is $2^{-51}$, because it is only after the $51^{\text {st }}$ card is dropped that we have no more choices. We can derive an exact expression for the entropy of the riffle shuffle in the following manner. Consider a deck of $N=2 K$ cards. The probability that we run out of choices after $K$ cards is the probability of the first $K$ cards dropped being all from one particular half-deck, which is $2 \cdot 2^{-K}$. Now let's ask what is the probability that we run out of choices after $(K+1)$ cards are dropped. If all the remaining ( $K-1$ ) cards are from the right half-deck, this means that we must have one of the R cards among the first $K$ dropped. Note that this R card cannot be the $(K+1)^{\text {th }}$ card dropped, since then all of the first $K$ cards are L, which we have already considered. Thus, there are $\binom{K}{1}=K$ such configurations, each with a probability $2^{-K-1}$. Next, suppose we run out of choices after $(K+2)$ cards are dropped. If the remaining ( $K-2$ ) cards are R , this means we must have 2 of the R cards among the first $(K+1)$ dropped, which means $\binom{K+1}{2}$ possibilities. Note that the $(K+2)^{\text {th }}$ card must be L , since if it were R this would mean that the last ( $K-1$ ) cards are R , which we have already considered. Continuing in this manner, we conclude

$$
\Omega_{K}=2 \sum_{n=0}^{K}\binom{K+n-1}{n}=\binom{2 K}{K}
$$

| $K$ | $\Omega_{K}$ | $S_{K}$ | $\log _{2}\binom{2 K}{K}$ |
| :---: | :---: | :---: | :---: |
| 2 | 6 | 2.500 | 2.585 |
| 12 | 2704156 | 20.132 | 20.367 |
| 26 | $4.96 \times 10^{14}$ | 46.274 | 48.817 |
| 100 | $9.05 \times 10^{58}$ | 188.730 | 195.851 |

Table 1: Riffle shuffle results.
and

$$
S_{K}=-\sum_{a=1}^{\Omega_{K}} p_{a} \log _{2} p_{a}=\sum_{n=0}^{K-1}\binom{K+n-1}{n} \cdot 2^{-(K+n)} \cdot(K+n)
$$

The results are tabulated below in Table 1. For a deck of 52 cards, the actual entropy per riffle shuffle is $S_{26}=$ 46.274 .
(d) Ignoring redundancies, we require $k=S_{\text {random }} /(\Delta S)_{\text {riffle }}=4.62$ shuffles if we assume all riffle outcomes are equally likely, and 4.88 if we use the exact result for the riffle entropy. Since there are no fractional shuffles, we round up to $k=5$ in both cases. In fact, computer experiments show that the answer is $k=9$. The reason we are so far off is that we have ignored redundancies, i.e. we have assumed that all the states produced by two consecutive riffle shuffles are distinct. They are not! For decks with asymptotically large numbers of cards $N \gg 1$, the number of riffle shuffles required is $k \simeq \frac{3}{2} \log _{2} N$. See D. Bayer and P. Diaconis, Annals of Applied Probability 2, 294 (1992).
(2) A six-sided die is loaded so that the probability to throw a three is twice that of throwing a two, and the probability of throwing a four is twice that of throwing a five.
(a) Find the distribution $\left\{p_{n}\right\}$ consistent with maximum entropy, given these constraints.
(b) Assuming the maximum entropy distribution, given two such identical dice, what is the probability to roll a total of seven if both are thrown simultaneously?

Solution :
(a) We have the following constraints:

$$
\begin{aligned}
& X^{0}(\boldsymbol{p})=p_{1}+p_{2}+p_{3}+p_{4}+p_{5}+p_{6}-1=0 \\
& X^{1}(\boldsymbol{p})=p_{3}-2 p_{2}=0 \\
& X^{2}(\boldsymbol{p})=p_{4}-2 p_{5}=0 .
\end{aligned}
$$

We define

$$
S^{*}(\boldsymbol{p}, \boldsymbol{\lambda}) \equiv-\sum_{n} p_{n} \ln p_{n}-\sum_{a=0}^{2} \lambda_{a} X^{(a)}(\boldsymbol{p}),
$$

and freely extremize over the probabilities $\left\{p_{1}, \ldots, p_{6}\right\}$ and the undetermined Lagrange multipliers $\left\{\lambda_{0}, \lambda_{1}, \lambda_{2}\right\}$. We obtain

$$
\begin{array}{ll}
\frac{\partial S^{*}}{\partial p_{1}}=-1-\ln p_{1}-\lambda_{0} & \frac{\partial S^{*}}{\partial p_{4}}=-1-\ln p_{4}-\lambda_{0}-\lambda_{2} \\
\frac{\partial S^{*}}{\partial p_{2}}=-1-\ln p_{2}-\lambda_{0}+2 \lambda_{1} & \frac{\partial S^{*}}{\partial p_{5}}=-1-\ln p_{5}-\lambda_{0}+2 \lambda_{2} \\
\frac{\partial S^{*}}{\partial p_{3}}=-1-\ln p_{3}-\lambda_{0}-\lambda_{1} & \frac{\partial S^{*}}{\partial p_{6}}=-1-\ln p_{6}-\lambda_{0} .
\end{array}
$$

Extremizing with respect to the undetermined multipliers generates the three constraint equations. We therefore have

$$
\begin{array}{ll}
p_{1}=e^{-\lambda_{0}-1} & p_{4}=e^{-\lambda_{0}-1} e^{-\lambda_{2}} \\
p_{2}=e^{-\lambda_{0}-1} e^{2 \lambda_{1}} & p_{5}=e^{-\lambda_{0}-1} e^{2 \lambda_{2}} \\
p_{3}=e^{-\lambda_{0}-1} e^{-\lambda_{1}} & p_{6}=e^{-\lambda_{0}-1}
\end{array}
$$

We solve for $\left\{\lambda_{0}, \lambda_{1}, \lambda_{2}\right\}$ by imposing the three constraints. Let $x \equiv p_{1}=p_{6}=e^{-\lambda_{0}-1}$. Then $p_{2}=x e^{2 \lambda_{1}}$, $p_{3}=x e^{-\lambda_{1}}, p_{4}=x e^{-\lambda_{2}}$, and $p_{5}=x e^{2 \lambda_{2}}$. We then have

$$
\begin{aligned}
& p_{3}=2 p_{2} \quad \Rightarrow \quad e^{-3 \lambda_{1}}=2 \\
& p_{4}=2 p_{5} \quad \Rightarrow \quad e^{-3 \lambda_{2}}=2 .
\end{aligned}
$$

We may now solve for $x$ :

$$
\sum_{n=1}^{6} p_{n}=\left(2+2^{1 / 3}+2^{4 / 3}\right) x=1 \quad \Rightarrow \quad x=\frac{1}{2+3 \cdot 2^{1 / 3}} .
$$

We now have all the probabilities:

$$
\begin{array}{ll}
p_{1}=x=0.1730 & p_{4}=2^{1 / 3} x=0.2180 \\
p_{2}=2^{-2 / 3} x=0.1090 & p_{5}=2^{-2 / 3} x=0.1090 \\
p_{3}=2^{1 / 3} x=0.2180 & p_{6}=x=0.1730 .
\end{array}
$$

(b) The probability to roll a seven with two of these dice is

$$
\begin{aligned}
P(7) & =2 p_{1} p_{6}+2 p_{2} p_{5}+2 p_{3} p_{4} \\
& =2\left(1+2^{-4 / 3}+2^{2 / 3}\right) x^{2}=0.1787 .
\end{aligned}
$$

(3) Consider the contraption in Fig. 2. At each of $k$ steps, a particle can fork to either the left ( $n_{j}=1$ ) or to the right $\left(n_{j}=0\right)$. The final location is then a $k$-digit binary number.
(a) Assume the probability for moving to the left is $p$ and the probability for moving to the right is $q \equiv 1-p$ at each fork, independent of what happens at any of the other forks. I.e. all the forks are uncorrelated. Compute $\left\langle X_{k}\right\rangle$. Hint: $X_{k}$ can be represented as a $k$-digit binary number, i.e. $X_{k}=n_{k-1} n_{k-2} \cdots n_{1} n_{0}=\sum_{j=0}^{k-1} 2^{j} n_{j}$.
(b) Compute $\left\langle X_{k}^{2}\right\rangle$ and the variance $\left\langle X_{k}^{2}\right\rangle-\left\langle X_{k}\right\rangle^{2}$.
(c) $X_{k}$ may be written as the sum of $k$ random numbers. Does $X_{k}$ satisfy the central limit theorem as $k \rightarrow \infty$ ? Why or why not?


Figure 2: Generator for a $k$-digit random binary number ( $k=4$ shown).
Solution :
(a) The position after $k$ forks can be written as a $k$-digit binary number: $n_{k-1} n_{k-2} \cdots n_{1} n_{0}$. Thus,

$$
X_{k}=\sum_{j=0}^{k-1} 2^{j} n_{j}
$$

where $n_{j}=0$ or 1 according to $P_{n}=p \delta_{n, 1}+q \delta_{n, 0}$. Now it is clear that $\left\langle n_{j}\right\rangle=p$, and therefore

$$
\left\langle X_{k}\right\rangle=p \sum_{j=0}^{k-1} 2^{j}=p \cdot\left(2^{k}-1\right)
$$

(b) The variance in $X_{k}$ is

$$
\begin{aligned}
\operatorname{Var}\left(X_{k}\right)=\left\langle X_{k}^{2}\right\rangle-\left\langle X_{k}\right\rangle^{2} & =\sum_{j=0}^{k-1} \sum_{j^{\prime}=0}^{k-1} 2^{j+j^{\prime}}\left(\left\langle n_{j} n_{j^{\prime}}\right\rangle-\left\langle n_{j}\right\rangle\left\langle n_{j^{\prime}}\right\rangle\right) \\
& =p(1-p) \sum_{j=0}^{k-1} 4^{j}=p(1-p) \cdot \frac{1}{3}\left(4^{k}-1\right)
\end{aligned}
$$

since $\left\langle n_{j} n_{j^{\prime}}\right\rangle-\left\langle n_{j}\right\rangle\left\langle n_{j^{\prime}}\right\rangle=p(1-p) \delta_{j j^{\prime}}$.
(c) Clearly the distribution of $X_{k}$ does not obey the CLT, since $\left\langle X_{k}\right\rangle$ scales exponentially with $k$. Also note

$$
\lim _{k \rightarrow \infty} \frac{\sqrt{\operatorname{Var}\left(X_{k}\right)}}{\left\langle X_{k}\right\rangle}=\sqrt{\frac{1-p}{3 p}}
$$

which is a constant. For distributions obeying the CLT, the ratio of the rms fluctuations to the mean scales as the inverse square root of the number of trials. The reason that this distribution does not obey the CLT is that the variance of the individual terms is increasing with $j$.
(4) The binomial distribution,

$$
B_{N}(n, p)=\binom{N}{n} p^{n}(1-p)^{N-n}
$$

tells us the probability for $n$ successes in $N$ trials if the individual trial success probability is $p$. The average number of successes is $\nu=\sum_{n=0}^{N} n B_{N}(n, p)=N p$. Consider the limit $N \rightarrow \infty$.
(a) Show that the probability of $n$ successes becomes a function of $n$ and $\nu$ alone. That is, evaluate

$$
P_{\nu}(n)=\lim _{N \rightarrow \infty} B_{N}(n, \nu / N)
$$

This is the Poisson distribution.
(b) Show that the moments of the Poisson distribution are given by

$$
\left\langle n^{k}\right\rangle=e^{-\nu}\left(\nu \frac{\partial}{\partial \nu}\right)^{k} e^{\nu}
$$

(c) Evaluate the mean and variance of the Poisson distribution.

The Poisson distribution is also known as the law of rare events since $p=\nu / N \rightarrow 0$ in the $N \rightarrow \infty$ limit. See http://en.wikipedia.org/wiki/Poisson_distribution\#Occurrence for some amusing applications of the Poisson distribution.

## Solution :

(a) We have

$$
P_{\nu}(n)=\lim _{N \rightarrow \infty} \frac{N!}{n!(N-n)!}\left(\frac{\nu}{N}\right)^{n}\left(1-\frac{\nu}{N}\right)^{N-n}
$$

Note that

$$
(N-n)!\simeq(N-n)^{N-n} e^{n-N}=N^{N-n}\left(1-\frac{n}{N}\right)^{N} e^{n-N} \rightarrow N^{N-n} e^{N}
$$

where we have used the result $\lim _{N \rightarrow \infty}\left(1+\frac{x}{N}\right)^{N}=e^{x}$. Thus, we find

$$
P_{\nu}(n)=\frac{1}{n!} \nu^{n} e^{-\nu},
$$

the Poisson distribution. Note that $\sum_{n=0}^{\infty} P_{n}(\nu)=1$ for any $\nu$.
(b) We have

$$
\begin{aligned}
\left\langle n^{k}\right\rangle & =\sum_{n=0}^{\infty} P_{\nu}(n) n^{k}=\sum_{n=0}^{\infty} \frac{1}{n!} n^{k} \nu^{n} e^{-\nu} \\
& =e^{-\nu}\left(\nu \frac{d}{d \nu}\right)^{k} \sum_{n=0}^{\infty} \frac{\nu^{n}}{n!}=e^{-\nu}\left(\nu \frac{\partial}{\partial \nu}\right)^{k} e^{\nu}
\end{aligned}
$$

(c) Using the result from (b), we have $\langle n\rangle=\nu$ and $\left\langle n^{2}\right\rangle=\nu+\nu^{2}$, hence $\operatorname{Var}(n)=\nu$.
(5) Let $P(x)=\left(2 \pi \sigma^{2}\right)^{-1 / 2} e^{-(x-\mu)^{2} / 2 \sigma^{2}}$. Compute the following integrals:
(a) $I=\int_{-\infty}^{\infty} d x P(x) x^{3}$.
(b) $I=\int_{-\infty}^{\infty} d x P(x) \cos (Q x)$.
(c) $I=\int_{-\infty}^{\infty} d x \int_{-\infty}^{\infty} d y P(x) P(y) e^{x y}$. You may set $\mu=0$ to make this somewhat simpler. Under what conditions does this expression converge?

Solution:
(a) Write

$$
x^{3}=(x-\mu+\mu)^{3}=(x-\mu)^{3}+3(x-\mu)^{2} \mu+3(x-\mu) \mu^{2}+\mu^{3},
$$

so that

$$
\left\langle x^{3}\right\rangle=\frac{1}{\sqrt{2 \pi \sigma^{2}}} \int_{-\infty}^{\infty} d t e^{-t^{2} / 2 \sigma^{2}}\left\{t^{3}+3 t^{2} \mu+3 t \mu^{2}+\mu^{3}\right\}
$$

Since $\exp \left(-t^{2} / 2 \sigma^{2}\right)$ is an even function of $t$, odd powers of $t$ integrate to zero. We have $\left\langle t^{2}\right\rangle=\sigma^{2}$, so

$$
\left\langle x^{3}\right\rangle=\mu^{3}+3 \mu \sigma^{2}
$$

A nice trick for evaluating $\left\langle t^{2 k}\right\rangle$ :

$$
\begin{aligned}
\left\langle t^{2 k}\right\rangle & =\frac{\int_{-\infty}^{\infty} d t e^{-\lambda t^{2}} t^{2 k}}{\int_{-\infty}^{\infty} d t e^{-\lambda t^{2}}}=\frac{(-1)^{k} \frac{d^{k}}{d \lambda^{k}} \int_{-\infty}^{\infty} d t e^{-\lambda t^{2}}}{\int_{-\infty}^{\infty} d t e^{-\lambda t^{2}}}=\left.\frac{(-1)^{k}}{\sqrt{\lambda}} \frac{d^{k} \sqrt{\lambda}}{d \lambda^{k}}\right|_{\lambda=1 / 2 \sigma^{2}} \\
& =\left.\frac{1}{2} \cdot \frac{3}{2} \cdots \frac{(2 k-1)}{2} \lambda^{-k}\right|_{\lambda=1 / 2 \sigma^{2}}=\frac{(2 k)!}{2^{k} k!} \sigma^{2 k} .
\end{aligned}
$$

(b) We have

$$
\begin{aligned}
\langle\cos (Q x)\rangle & =\operatorname{Re}\left\langle e^{i Q x}\right\rangle=\operatorname{Re}\left[\frac{e^{i Q \mu}}{\sqrt{2 \pi \sigma^{2}}} \int_{-\infty}^{\infty} d t e^{-t^{2} / 2 \sigma^{2}} e^{i Q t}\right] \\
& =\operatorname{Re}\left[e^{i Q \mu} e^{-Q^{2} \sigma^{2} / 2}\right]=\cos (Q \mu) e^{-Q^{2} \sigma^{2} / 2}
\end{aligned}
$$

Here we have used the result

$$
\frac{1}{\sqrt{2 \pi \sigma^{2}}} \int_{-\infty}^{\infty} d t e^{-\alpha t^{2}-\beta t}=\sqrt{\frac{\pi}{\alpha}} e^{\beta^{2} / 4 \alpha}
$$

with $\alpha=1 / 2 \sigma^{2}$ and $\beta=-i Q$. Another way to do it is to use the general result derive above in part (a) for $\left\langle t^{2 k}\right\rangle$ and do the sum:

$$
\begin{aligned}
\langle\cos (Q x)\rangle & =\operatorname{Re}\left\langle e^{i Q x}\right\rangle=\operatorname{Re}\left[\frac{e^{i Q \mu}}{\sqrt{2 \pi \sigma^{2}}} \int_{-\infty}^{\infty} d t e^{-t^{2} / 2 \sigma^{2}} e^{i Q t}\right] \\
& =\cos (Q \mu) \sum_{k=0}^{\infty} \frac{\left(-Q^{2}\right)^{k}}{(2 k)!}\left\langle t^{2 k}\right\rangle=\cos (Q \mu) \sum_{k=0}^{\infty} \frac{1}{k!}\left(-\frac{1}{2} Q^{2} \sigma^{2}\right)^{k} \\
& =\cos (Q \mu) e^{-Q^{2} \sigma^{2} / 2}
\end{aligned}
$$

(c) We have

$$
I=\int_{-\infty}^{\infty} d x \int_{-\infty}^{\infty} d y P(x) P(y) e^{\kappa^{2} x y}=\frac{e^{-\mu^{2} / 2 \sigma^{2}}}{2 \pi \sigma^{2}} \int d^{2} x e^{-\frac{1}{2} A_{i j} x_{i} x_{j}} e^{b_{i} x_{i}},
$$

where $\boldsymbol{x}=(x, y)$,

$$
A=\left(\begin{array}{cc}
\sigma^{2} & -\kappa^{2} \\
-\kappa^{2} & \sigma^{2}
\end{array}\right) \quad, \quad \boldsymbol{b}=\binom{\mu / \sigma^{2}}{\mu / \sigma^{2}} .
$$

Using the general formula for the Gaussian integral,

$$
\int d^{n} x e^{-\frac{1}{2} A_{i j} x_{i} x_{j}} e^{b_{i} x_{i}}=\frac{(2 \pi)^{n / 2}}{\sqrt{\operatorname{det}(A)}} \exp \left(\frac{1}{2} A_{i j}^{-1} b_{i} b_{j}\right)
$$

we obtain

$$
I=\frac{1}{\sqrt{1-\kappa^{4} \sigma^{4}}} \exp \left(\frac{\mu^{2} \kappa^{2}}{1-\kappa^{2} \sigma^{2}}\right) .
$$

Convergence requires $\kappa^{2} \sigma^{2}<1$.
(6) A rare disease is known to occur in $f=0.02 \%$ of the general population. Doctors have designed a test for the disease with $\nu=99.90 \%$ sensitivity and $\rho=99.95 \%$ specificity.
(a) What is the probability that someone who tests positive for the disease is actually sick?
(b) Suppose the test is administered twice, and the results of the two tests are independent. If a random individual tests positive both times, what are the chances he or she actually has the disease?
(c) For a binary partition of events, find an expression for $P(X \mid A \cap B)$ in terms of $P(A \mid X), P(B \mid X), P(A \mid \neg X)$, $P(B \mid \neg X)$, and the priors $P(X)$ and $P(\neg X)=1-P(X)$. You should assume $A$ and $B$ are independent, so $P(A \cap B \mid X)=P(A \mid X) \cdot P(B \mid X)$.

## Solution :

(a) Let $X$ indicate that a person is infected, and $A$ indicate that a person has tested positive. We then have $\nu=$ $P(A \mid X)=0.9990$ is the sensitivity and $\rho=P(\neg A \mid \neg X)=0.9995$ is the specificity. From Bayes' theorem, we have

$$
P(X \mid A)=\frac{P(A \mid X) \cdot P(X)}{P(A \mid X) \cdot P(X)+P(A \mid \neg X) \cdot P(\neg X)}=\frac{\nu f}{\nu f+(1-\rho)(1-f)},
$$

where $P(A \mid \neg X)=1-P(\neg A \mid \neg X)=1-\rho$ and $P(X)=f$ is the fraction of infected individuals in the general population. With $f=0.0002$, we find $P(X \mid A)=0.2856$.
(b) We now need

$$
P\left(X \mid A^{2}\right)=\frac{P\left(A^{2} \mid X\right) \cdot P(X)}{P\left(A^{2} \mid X\right) \cdot P(X)+P\left(A^{2} \mid \neg X\right) \cdot P(\neg X)}=\frac{\nu^{2} f}{\nu^{2} f+(1-\rho)^{2}(1-f)},
$$

where $A^{2}$ indicates two successive, independent tests. We find $P\left(X \mid A^{2}\right)=0.9987$.
(c) Assuming $A$ and $B$ are independent, we have

$$
\begin{aligned}
P(X \mid A \cap B) & =\frac{P(A \cap B \mid X) \cdot P(X)}{P(A \cap B \mid X) \cdot P(X)+P(A \cap B \mid \neg X) \cdot P(\neg X)} \\
& =\frac{P(A \mid X) \cdot P(B \mid X) \cdot P(X)}{P(A \mid X) \cdot P(B \mid X) \cdot P(X)+P(A \mid \neg X) \cdot P(B \mid \neg X) \cdot P(\neg X)} .
\end{aligned}
$$

This is exactly the formula used in part (b).
(7) Let $p(x)=\operatorname{Pr}[X=x]$ where $X$ is a discrete random variable and both $X$ and $x$ are taken from an 'alphabet' $\mathcal{X}$. Let $p(x, y)$ be a normalized joint probability distribution on two random variables $X \in \mathcal{X}$ and $Y \in \mathcal{Y}$. The entropy of the joint distribution is $S(X, Y)=-\sum_{x, y} p(x, y) \log p(x, y)$. The conditional probability $p(y \mid x)$ for $y$ given $x$ is defined as $p(y \mid x)=p(x, y) / p(x)$, where $p(x)=\sum_{y} p(x, y)$.
(a) Show that the conditional entropy $S(Y \mid X)=-\sum_{x, y} p(x, y) \log p(y \mid x)$ satisfies

$$
S(Y \mid X)=S(X, Y)-S(X)
$$

Thus, conditioning reduces the entropy, and the entropy of a pair of random variables is the sum of the entropy of one plus the conditional entropy of the other.
(b) The mutual information $I(X, Y)$ is

$$
I(X, Y)=\sum_{x, y} p(x, y) \log \left(\frac{p(x, y)}{p(x) p(y)}\right)
$$

Show that

$$
I(X, Y)=S(X)+S(Y)-S(X, Y)
$$

(c) Show that $S(Y \mid X) \leq S(Y)$ and that the equality holds only when $X$ and $Y$ are independently distributed.

Solution :
(a) We have

$$
\begin{aligned}
S(Y \mid X) & =-\sum_{x, y} p(x, y) \log \left(\frac{p(x, y)}{p(x)}\right) \\
& =-\sum_{x, y} p(x, y) \log p(x, y)+\sum_{x, y} p(x, y) \log p(x) \\
& =S(X, Y)-S(X)
\end{aligned}
$$

since $\sum_{y} p(x, y)=p(x)$.
(b) Clearly

$$
\begin{aligned}
I(X, Y) & =\sum_{x, y} p(x, y)(\log p(x, y)-\log p(x)-\log p(y)) \\
& =\sum_{x, y} p(x, y) \log p(x, y)-\sum_{x} p(x) \log p(x)-\sum_{y} p(y) \log p(y) \\
& =S(X)+S(Y)-S(X, Y)
\end{aligned}
$$

(c) Let's introduce some standard notation ${ }^{1}$. Denote the average of a function of a random variable as

$$
\mathbb{E} f(X)=\sum_{x \in \mathcal{X}} p(x) f(x)
$$

Equivalently, we could use the familiar angular bracket notation for averages and write $\mathbb{E} f(X)$ as $\langle f(X)\rangle$. For averages over several random variables, we have

$$
\mathbb{E} f(X, Y)=\sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} p(x, y) f(x, y)
$$

[^0]et cetera. Now here's a useful fact. If $f(x)$ is a convex function, then
$$
\mathbb{E} f(X) \geq f(\mathbb{E} X)
$$

For continuous functions, $f(x)$ is convex if $f^{\prime \prime}(x) \geq 0$ everywhere $^{2}$. If $f(x)$ is convex on some interval $[a, b]$, then for $x_{1,2} \in[a, b]$ we must have

$$
f\left(\lambda x_{1}+(1-\lambda) x_{2}\right) \leq \lambda f\left(x_{1}\right)+(1-\lambda) f\left(x_{2}\right)
$$

This is easily generalized to

$$
f\left(\sum_{n} p_{n} x_{n}\right) \leq \sum_{n} p_{n} f\left(x_{n}\right)
$$

where $\sum_{n} p_{n}=1$, which proves our claim - a result known as Jensen's theorem.
Now $\log x$ is concave, hence the function $f(x)=-\log x$ is convex, and therefore

$$
\begin{aligned}
I(X, Y) & =-\sum_{x, y} p(x, y) \log \left(\frac{p(x) p(y)}{p(x, y)}\right)=-\mathbb{E} \log \left(\frac{p(X) p(Y)}{p(X, Y)}\right) \\
& \geq-\log \mathbb{E}\left(\frac{p(X) p(Y)}{p(X, Y)}\right)=-\log \left(\sum_{x, y} p(x, y) \cdot \frac{p(x) p(y)}{p(x, y)}\right) \\
& =-\log \left(\sum_{x} p(x) \cdot \sum_{y} p(y)\right)=-\log (1)=0
\end{aligned}
$$

So $I(X, Y) \geq 0$. Clearly $I(X, Y)=0$ when $X$ and $Y$ are independently distributed, i.e. when $p(x, y)=p(x) p(y)$. Using the results from part (b), we then have

$$
\begin{aligned}
I(X, Y) & =S(X)+S(Y)-S(X, Y) \\
& =S(Y)-S(Y \mid X) \geq 0
\end{aligned}
$$

[^1](8) The $n^{\text {th }}$ moment of the normalized Gaussian distribution $P(x)=(2 \pi)^{-1 / 2} \exp \left(-\frac{1}{2} x^{2}\right)$ is defined by
$$
\left\langle x^{n}\right\rangle=\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{\infty} d x x^{n} \exp \left(-\frac{1}{2} x^{2}\right)
$$

Clearly $\left\langle x^{n}\right\rangle=0$ if $n$ is a nonnegative odd integer. Next consider the generating function

$$
Z(j)=\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{\infty} d x \exp \left(-\frac{1}{2} x^{2}\right) \exp (j x)=\exp \left(\frac{1}{2} j^{2}\right)
$$

(a) Show that

$$
\left\langle x^{n}\right\rangle=\left.\frac{d^{n} Z}{d j^{n}}\right|_{j=0}
$$

and provide an explicit result for $\left\langle x^{2 k}\right\rangle$ where $k \in \mathbb{N}$.
(b) Now consider the following integral:

$$
F(\lambda)=\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{\infty} d x \exp \left(-\frac{1}{2} x^{2}-\frac{\lambda}{4!} x^{4}\right)
$$

The integral has no known analytic form ${ }^{3}$, but we may express the result as a power series in the parameter $\lambda$ by Taylor expanding $\exp \left(-\frac{\lambda}{4!} x^{4}\right)$ and then using the result of part (a) for the moments $\left\langle x^{4 k}\right\rangle$. Find the coefficients in the perturbation expansion,

$$
F(\lambda)=\sum_{k=0}^{\infty} C_{k} \lambda^{k}
$$

(c) Define the remainder after $N$ terms as

$$
R_{N}(\lambda)=F(\lambda)-\sum_{k=0}^{N} C_{k} \lambda^{k}
$$

Compute $R_{N}(\lambda)$ by evaluating numerically the integral for $F(\lambda)$ (using Mathematica or some other numerical package) and subtracting the finite sum. Then define the ratio $S_{N}(\lambda)=R_{N}(\lambda) / F(\lambda)$, which is the relative error from the $N$ term approximation and plot the absolute relative error $\left|S_{N}(\lambda)\right|$ versus $N$ for several values of $\lambda$.(I suggest you plot the error on a log scale.) What do you find?? Try a few values of $\lambda$ including $\lambda=0.01$, $\lambda=0.05, \lambda=0.2, \lambda=0.5, \lambda=1, \lambda=2$.
(d) Repeat the calculation for the integral

$$
G(\lambda)=\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{\infty} d x \exp \left(-\frac{1}{2} x^{2}-\frac{\lambda}{6!} x^{6}\right)
$$

(e) Reflect meaningfully on the consequences for weakly and strongly coupled quantum field theories.

[^2]Solution :
(a) Clearly

$$
\left.\frac{d^{n}}{d j^{n}}\right|_{j=0} e^{j x}=x^{n}
$$

so $\left\langle x^{n}\right\rangle=\left(d^{n} Z / d j^{n}\right)_{j=0}$. With $Z(j)=\exp \left(\frac{1}{2} j^{2}\right)$, only the $k^{\text {th }}$ order term in $j^{2}$ in the Taylor series for $Z(j)$ contributes, and we obtain

$$
\left\langle x^{2 k}\right\rangle=\frac{d^{2 k}}{d j^{2 k}}\left(\frac{j^{2 k}}{2^{k} k!}\right)=\frac{(2 k)!}{2^{k} k!}
$$

(b) We have

$$
F(\lambda)=\sum_{n=0}^{\infty} \frac{1}{n!}\left(-\frac{\lambda}{4!}\right)^{n}\left\langle x^{4 n}\right\rangle=\sum_{n=0}^{\infty} \frac{(4 n)!}{4^{n}(4!)^{n} n!(2 n)!}(-\lambda)^{n}
$$

This series is asymptotic. It has the properties

$$
\lim _{\lambda \rightarrow 0} \frac{R_{N}(\lambda)}{\lambda^{N}}=0 \quad(\text { fixed } N) \quad, \quad \lim _{N \rightarrow \infty} \frac{R_{N}(\lambda)}{\lambda^{N}}=\infty \quad(\text { fixed } \lambda)
$$

where $R_{N}(\lambda)$ is the remainder after $N$ terms, defined in part (c). The radius of convergence is zero. To see this, note that if we reverse the sign of $\lambda$, then the integrand of $F(\lambda)$ diverges badly as $x \rightarrow \pm \infty$. So $F(\lambda)$ is infinite for $\lambda<0$, which means that there is no disk of any finite radius of convergence which encloses the point $\lambda=0$. Note that by Stirling's rule,

$$
(-1)^{n} C_{n} \equiv \frac{(4 n)!}{4^{n}(4!)^{n} n!(2 n)!} \sim n^{n} \cdot\left(\frac{2}{3}\right)^{n} e^{-n} \cdot(\pi n)^{-1 / 2}
$$

and we conclude that the magnitude of the summand reaches a minimum value when $n=n^{*}(\lambda)$, with

$$
n^{*}(\lambda) \approx \frac{3}{2 \lambda}
$$

for small values of $\lambda$. For large $n$, the magnitude of the coefficient $C_{n}$ grows as $\left|C_{n}\right| \sim e^{n \ln n+\mathcal{O}(n)}$, which dominates the $\lambda^{n}$ term, no matter how small $\lambda$ is.
(c) Results are plotted in fig. 3.

It is worth pointing out that the series for $F(\lambda)$ and for $\ln F(\lambda)$ have diagrammatic interpretations. For a Gaussian integral, one has

$$
\left\langle x^{2 k}\right\rangle=\left\langle x^{2}\right\rangle^{k} \cdot A_{2 k}
$$

where $A_{2 k}$ is the number of contractions. For a proof, see $\S 1.4 .3$ of the notes. For our integral, $\left\langle x^{2}\right\rangle=1$. The number of contractions $A_{2 k}$ is computed in the following way. For each of the $2 k$ powers of $x$, we assign an index running from 1 to $2 k$. The indices are contracted, i.e. paired, with each other. How many pairings are there? Suppose we start with any from among the $2 k$ indices. Then there are $(2 k-1)$ choices for its mate. We then choose another index arbitrarily. There are now $(2 k-3)$ choices for $i t s$ mate. Carrying this out to its completion, we find that the number of contractions is

$$
A_{2 k}=(2 k-1)(2 k-3) \cdots 3 \cdot 1=\frac{(2 k)!}{2^{k} k!}
$$

exactly as we found in part (a). Now consider the integral $F(\lambda)$. If we expand the quartic term in a power series, then each power of $\lambda$ brings an additional four powers of $x$. It is therefore convenient to represent each such quartet with the symbol $\times$. At order $N$ of the series expansion, we have $N \times$ 's and $4 N$ indices to contract. Each full contraction of the indices may be represented as a labeled diagram, which is in general composed of several disjoint connected subdiagrams. Let us label these subdiagrams, which we will call clusters, by an index $\gamma$. Now


Figure 3: Relative error versus number of terms kept for the asymptotic series for $F(\lambda)$. Note that the optimal number of terms to sum is $N^{*}(\lambda) \approx \frac{3}{2 \lambda}$.
suppose we have a diagram consisting of $m_{\gamma}$ subdiagrams of type $\gamma$, for each $\gamma$. If the cluster $\gamma$ contains $n_{\gamma}$ vertices $(\times)$, then we must have

$$
N=\sum_{\gamma} m_{\gamma} n_{\gamma}
$$

How many ways are there of assigning the labels to such a diagram? One might think $(4!)^{N} \cdot N$ !, since for each vertex $\times$ there are 4 ! permutations of its four labels, and there are $N$ ! ways to permute all the vertices. However, this overcounts diagrams which are invariant under one or more of these permutations. We define the symmetry factor $s_{\gamma}$ of the (unlabeled) cluster $\gamma$ as the number of permutations of the indices of a corresponding labeled cluster which result in the same contraction. We can also permute the $m_{\gamma}$ identical disjoint clusters of type $\gamma$.

Examples of clusters and their corresponding symmetry factors are provided in fig. 4, for all diagrams with $n_{\gamma} \leq 3$. There is only one diagram with $n_{\gamma}=1$, resembling @. To obtain $s_{\gamma}=8$, note that each of the circles can be separately rotated by an angle $\pi$ about the long symmetry axis. In addition, the figure can undergo a planar rotation by $\pi$ about an axis which runs through the sole vertex and is normal to the plane of the diagram. This results in $s_{\gamma}=2 \cdot 2 \cdot 2=8$. For the cluster COO, there is one extra circle, so $s_{\gamma}=2^{4}=16$. The third diagram in figure shows two vertices connected by four lines. Any of the 4 ! permutations of these lines results in the same diagram. In addition, we may reflect about the vertical symmetry axis, interchanging the vertices, to obtain another symmetry operation. Thus $s_{\gamma}=2 \cdot 4!=48$. One might ask why we don't also count the planar rotation by $\pi$ as a symmetry operation. The answer is that it is equivalent to a combination of a reflection and a permutation, so it is not in fact a distinct symmetry operation. (If it were distinct, then $s_{\gamma}$ would be 96.) Finally, consider the last diagram in the figure, which resembles a sausage with three links joined at the ends into a circle. If we keep the vertices fixed, there are 8 symmetry operations associated with the freedom to exchange the two lines associated with each of the three sausages. There are an additional 6 symmetry operations associated with permuting the three vertices, which can be classified as three in-plane rotations by $0, \frac{2 \pi}{3}$ and $\frac{4 \pi}{3}$, each of which can also be combined with a reflection about the $y$-axis (this is known as the group $C_{3 v}$ ). Thus, $s_{\gamma}=8 \cdot 6=48$.

Now let us compute an expression for $F(\gamma)$ in terms of the clusters. We sum over all possible numbers of clusters


Figure 4: Cluster symmetry factors. A vertex is represented as a black dot $(\bullet)$ with four 'legs'.
at each order:

$$
\begin{aligned}
F(\gamma) & =\sum_{N=0}^{\infty} \frac{1}{N!} \sum_{\left\{m_{\gamma}\right\}} \frac{(4!)^{N} N!}{\prod_{\gamma} s_{\gamma}^{m_{\gamma}} m_{\gamma}!}\left(-\frac{\lambda}{4!}\right)^{N} \delta_{N, \sum_{\gamma} m_{\gamma} n_{\gamma}} \\
& =\exp \left(\sum_{\gamma} \frac{(-\lambda)^{n_{\gamma}}}{s_{\gamma}}\right) .
\end{aligned}
$$

Thus,

$$
\ln F(\gamma)=\sum_{\gamma} \frac{(-\lambda)^{n_{\gamma}}}{s_{\gamma}}
$$

and the logarithm of the sum over all diagrams is a sum over connected clusters. It is instructive to work this out to order $\lambda^{2}$. We have, from the results of part (b),

$$
F(\lambda)=1-\frac{1}{8} \lambda+\frac{35}{384} \lambda^{2}+\mathcal{O}\left(\lambda^{3}\right) \quad \Longrightarrow \quad \ln F(\lambda)=-\frac{1}{8} \lambda+\frac{1}{12} \lambda^{2}+\mathcal{O}\left(\lambda^{3}\right)
$$

Note that there is one diagram with $N=1$ vertex, with symmetry factor $s=8$. For $N=2$ vertices, there are two diagrams, one with $s=16$ and one with $s=48$ (see fig. 4). Since $\frac{1}{16}+\frac{1}{48}=\frac{1}{12}$, the diagrammatic expansion is verified to order $\lambda^{2}$.
(d) We now have ${ }^{4}$

$$
\begin{aligned}
G(\lambda) & =\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{\infty} d x \exp \left(-\frac{1}{2} x^{2}-\frac{\lambda}{6!} x^{6}\right) \\
& =\sum_{n=0}^{\infty} \frac{1}{n!}\left(-\frac{\lambda}{6!}\right)^{n}\left\langle x^{6 n}\right\rangle=\sum_{n=0}^{\infty} C_{n} \lambda^{n},
\end{aligned}
$$

where

$$
C_{n}=\frac{(-1)^{n}(6 n)!}{(6!)^{n} n!2^{3 n}(3 n)!} .
$$

Invoking Stirling's approximation, we find

$$
\ln \left|C_{n}\right| \sim 2 n \ln n-\left(2+\ln \frac{5}{3}\right) n .
$$

[^3]

Figure 5: Logarithm of ratio of remainder after $N$ terms $R_{N}(\lambda)$ to the value of the integral $G(\lambda)$, for various values of $\lambda$.

From the above expression for $C_{n}$, we see that the magnitude of the contribution of the $n^{\text {th }}$ term in the perturbation series is

$$
C_{n} \lambda^{n}=(-1)^{n} \exp \left(2 n \ln n-\left(2+\ln \frac{10}{3}\right) n+n \ln \lambda\right)
$$

Differentiating, we find that this contribution is minimized for $n=n^{*}(\lambda)$, where

$$
n^{*}(\lambda)=\sqrt{\frac{10}{3 \lambda}}
$$

Via numerical integration using FORTRAN subroutines from QUADPACK, one obtains the results in Fig. 5 and Tab. 2.

| $\lambda$ | 10 | 2 | 0.5 | 0.2 | 0.1 | 0.05 | 0.02 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $F$ | 0.92344230 | 0.97298847 | 0.99119383 | 0.996153156 | 0.99800488 | 0.99898172 | 0.99958723 |
| $n^{*}$ | 0.68 | 1.3 | 2.6 | 4.1 | 5.8 | 8.2 | 13 |

Table 2: $F(\lambda)$ and $n^{*}(\lambda)$ for problem 8 d .

The series for $G(\lambda)$ and for $\ln G(\lambda)$ again have diagrammatic interpretations. If we expand the sextic term in a power series, each power of $\lambda$ brings an additional six powers of $x$. It is natural to represent each such sextet with as a vertex with six legs. At order $N$ of the series expansion, we have $N$ such vertices and $6 N$ legs to contract. As before, each full contraction of the leg indices may be represented as a labeled diagram, which is in general composed of several disjoint connected clusters. If the cluster $\gamma$ contains $n_{\gamma}$ vertices, then for any diagram we again must have $N=\sum_{\gamma} m_{\gamma} n_{\gamma}$, where $m_{\gamma}$ is the number of times the cluster $\gamma$ appears. As with the quartic example, the number of ways of assigning labels to a given diagram is given by the total number of possible permutations $(6!)^{N} \cdot N$ ! divided by a correction factor $\prod_{\gamma} s_{\gamma}^{m} m_{\gamma}$ !, where $s_{\gamma}$ is the symmetry factor of the cluster $\gamma$, and the $m_{\gamma}$ ! term accounts for the possibility of permuting among different labeled clusters of the same type $\gamma$.


Figure 6: Diagrams and their symmetry factors for the $\frac{1}{6!} \lambda x^{6}$ zero-dimensional field theory.

Examples of clusters and their corresponding symmetry factors are provided in Fig. 6. There is only one diagram with $n_{\gamma}=1$, shown panel (a), resembling a three-petaled flower. To obtain $s_{\gamma}=48$, note that each of the petals can be rotated by $180^{\circ}$ about an axis bisecting the petal, yielding a factor of $2^{3}$. The three petals can then be permuted, yielding an additional factor of 3!. Hence the total symmetry factor is $s_{\gamma}=2^{3} \cdot 3!=48$. Now we can see how dividing by the symmetry factor saves us from overcounting. In this case, we get $6!/ s_{\gamma}=720 / 48=$ $15=5 \cdot 3 \cdot 1$, which is the correct number of contractions. For the diagram in panel (b), the four petals and the central loop can each be rotated about a symmetry axis, yielding a factor $2^{5}$. The two left petals can be permuted, as can the two right petals. Finally, the two vertices can themselves be permuted. Thus, the symmetry factor is $s_{\gamma}=2^{5} \cdot 2^{2} \cdot 2=2^{8}=256$. In panel (c), the six lines can be permuted (6!) and the vertices can be exchanged (2), hence $s_{\gamma}=6!\cdot 2=1440$. In panel (d), the two outer loops each can be twisted by $180^{\circ}$, the central four lines can be permuted, and the vertices can be permuted, hence $s_{\gamma}=2^{2} \cdot 4!\cdot 2=192$. Finally, in panel (e), each pair of vertices is connected by three lines which can be permuted, and the vertices themselves can be permuted, so $s_{\gamma}=(3!)^{3} \cdot 3!=1296$.

Now let us compute an expression for $F(\gamma)$ in terms of the clusters. We sum over all possible numbers of clusters at each order:

$$
\begin{aligned}
G(\gamma) & =\sum_{N=0}^{\infty} \frac{1}{N!} \sum_{\left\{m_{\gamma}\right\}} \frac{(6!)^{N} N!}{\prod_{\gamma} s_{\gamma}^{m_{\gamma}} m_{\gamma}!}\left(-\frac{\lambda}{6!}\right)^{N} \delta_{N, \Sigma_{\gamma} m_{\gamma} n_{\gamma}} \\
& =\exp \left(\sum_{\gamma} \frac{(-\lambda)^{n_{\gamma}}}{s_{\gamma}}\right)
\end{aligned}
$$

Thus,

$$
\ln G(\gamma)=\sum_{\gamma} \frac{(-\lambda)^{n_{\gamma}}}{s_{\gamma}}
$$

and the logarithm of the sum over all diagrams is a sum over connected clusters. It is instructive to work this out to order $\lambda^{2}$. We have, from the results of part (a),

$$
G(\lambda)=1-\frac{\lambda}{2^{6} \cdot 3}+\frac{7 \cdot 11 \cdot \lambda^{2}}{2^{9} \cdot 3 \cdot 5}+\mathcal{O}\left(\lambda^{3}\right) \quad \Longrightarrow \quad \ln G(\lambda)=-\frac{\lambda}{2^{6} \cdot 3}+\frac{113 \cdot \lambda^{2}}{2^{8} \cdot 3^{2} \cdot 5}+\mathcal{O}\left(\lambda^{3}\right)
$$

Note that there is one diagram with $N=1$ vertex, with symmetry factor $s=48$. For $N=2$ vertices, there are three diagrams, one with $s=256$, one with $s=1440$, and one with $s=192$ (see Fig. 6). Since $\frac{1}{256}+\frac{1}{1440}+\frac{1}{192}=\frac{113}{2^{8} 3^{2} 5}$, the diagrammatic expansion is verified to order $\lambda^{2}$.
(e) In quantum field theory (QFT), the vertices themselves carry space-time labels, and the contractions, i.e. the lines connecting the legs of the vertices, are propagators $G\left(x_{i}^{\mu}-x_{j}^{\mu}\right)$, where $x_{i}^{\mu}$ is the space-time label associated with vertex $i$. It is convenient to work in momentum-frequency space, in which case we work with the Fourier transform $\hat{G}\left(p^{\mu}\right)$ of the space-time propagators. Integrating over the space-time coordinates of each vertex then enforces total 4-momentum conservation at each vertex. We then must integrate over all the internal 4-momenta to obtain the numerical value for a given diagram. The diagrams, as you know, are associated with Feynman's approach to QFT and are known as Feynman diagrams. Our example here is equivalent to a $(0+0)$-dimensional field theory, i.e. zero space dimensions and zero time dimensions. There are then no internal 4-momenta to integrate over, and each propagator is simply a number rather than a function. The discussion above of symmetry factors $s_{\gamma}$ carries over to the more general QFT case.

There is an important lesson to be learned here about the behavior of asymptotic series. As we have seen, if $\lambda$ is sufficiently small, summing more and more terms in the perturbation series results in better and better results, until one reaches an optimal order when the error is minimized. Beyond this point, summing additional terms makes the result worse, and indeed the perturbation series diverges badly as $N \rightarrow \infty$. Typically the optimal order of perturbation theory is inversely proportional to the coupling constant. For quantum electrodynamics (QED), where the coupling constant is the fine structure constant $\alpha=e^{2} / \hbar c \approx \frac{1}{137}$, we lose the ability to calculate in a reasonable time long before we get to 137 loops, so practically speaking no problems arise from the lack of convergence. In quantum chromodynamics (QCD), however, the effective coupling constant is about two orders of magnitude larger, and perturbation theory is a much more subtle affair.


[^0]:    ${ }^{1}$ See e.g. T. M. Cover and J. A. Thomas, Elements of Information Theory, 2nd edition (Wiley, 2006).

[^1]:    ${ }^{2} \mathrm{~A}$ concave function $g(x)$ is one for which $f(x)=-g(x)$ is convex.

[^2]:    ${ }^{3}$ In fact, it does. According to Mathematica, $F(\lambda)=\sqrt{\frac{2 u}{\pi}} \exp (u) K_{1 / 4}(u)$, where $u=3 / 4 \lambda$ and $K_{\nu}(z)$ is the modified Bessel function. I am grateful to Prof. John McGreevy for pointing this out.

[^3]:    ${ }^{4}$ According to Mathematica, the $G(\lambda)$ has the analytic form $G(\lambda)=\pi \sqrt{u}\left[\mathrm{Ai}^{2}(u)+\mathrm{Bi}^{2}(u)\right]$, where $u=(15 / 2 \lambda)^{1 / 3}$ and $\mathrm{Ai}(z)$ and $\mathrm{Bi}(z)$ are Airy functions. The definitions and properties of the Airy functions are discussed in $\S 9.2$ of the NIST Handbook of Mathematical Functions.

